

19 approach considered here does not rely on the proximity of approximate and resolved
20 models and can employ much coarser and inexpensive models to guide the fine-scale
21 simulations. Numerical results for three-phase flow and transport demonstrate the ad-
22 vantages, efficiency and utility of the method for uncertainty assessment in the history
23 matching.

24 **1 Introduction**

25 Uncertainties on the detailed description of reservoir lithofacies, porosity, and permeability
26 are major contributors to uncertainty in reservoir performance forecasting. Reducing this
27 uncertainty can be achieved by integrating additional data in subsurface modeling. With the
28 increasing interest in accurate prediction of subsurface properties, subsurface characteriza-
29 tion based on dynamic data, such as production data including gas/oil ratio becomes more
30 important.

31 To predict future reservoir performance, the reservoir properties, such as porosity and
32 permeability, need to be conditioned to production data. It is essential that the permeability
33 (and porosity) realizations adequately reflect the uncertainty in the reservoir properties, i.e.,
34 the probability distribution is sampled correctly. The uncertainty quantification is typically
35 carried out in a Bayesian framework where multiple realizations are sampled from a posterior
36 distribution that incorporates the prior information (e.g., [18]). This problem is challenging
37 because the permeability field is a function defined on a large number of grid blocks and
38 the production data nonlinearly depends on permeability. The Markov Chain Monte Carlo
39 (MCMC) method and its modifications have been used in previous findings to sample the
40 posterior distribution (e.g., [24]).

41 The direct MCMC simulations are generally very CPU demanding because each proposal
42 requires solving a forward coupled nonlinear partial differential equations over a large time
43 interval. The forward fine-scale problem is usually formulated on a large number of grid

44 blocks, which makes it prohibitively expensive to perform sufficient number of MCMC sim-
45 ulations. There have been a few attempts to propose MCMC methods with high acceptance
46 rate (e.g., [18, 24, 25]). For example, the randomized maximum likelihood method uses
47 unconditional realizations of the production and permeability data and solves a determin-
48 istic gradient-based inverse problem. The solution of this minimization problem is taken
49 as a proposal, and is accepted with probability one, because the rigorous acceptance prob-
50 ability is very difficult to estimate. Though efficient in many cases, this method may not
51 properly sample the posterior distribution [22]. Thus, developing efficient rigorous MCMC
52 calculations with high acceptance rate remains a challenging problem.

53 In this paper, we extend two-stage MCMC methods considered before (e.g., [10, 11]).
54 The two-stage MCMC involves a pre-screening stage where the proposals are screened using
55 approximate models. If the proposal is accepted in the first stage (screening stage), then the
56 resolved computations are performed to compute the acceptance probability. The novelty
57 of the proposed approach is two-fold. First, we employ error modeling ([16, 28, 15, 4,
58 26, 20]) which allows mapping the coarse-scale data to the fine-scale data via a nonlinear
59 map. The main goal of the error modeling is to construct a map from the coarse-scale
60 errors between computed and observed data to the fine-scale errors based on some prior
61 (offline) computations. The mapping between these low dimensional quantities is often
62 easy to construct based on fewer samples. Secondly, we consider inexpensive ensemble level
63 upscaling type methods for coarse-scale modeling (cf. [2]). To our best knowledge the error
64 models have not yet been used in rigorous sampling methods. Previous approaches within
65 two-stage MCMC methods [11, 10] have used only linear relation between coarse- and fine-
66 scale responses. Though it is found to be effective in many cases, linear relations may not be
67 very reliable when coarse-scale models both across spatial scales and uncertainties (in highly
68 nonlinear equations) are considered. In particular, for the examples presented here we have
69 found poor performance when linear models are used.

70 Ensemble level upscaling methods compute upscaled quantities which represent not only
71 subgrid variations, but also variations across the ensemble. These approaches compute
72 upscaled coefficients based on some sampled realizations. In this paper, we use some se-
73 lected realizations based on the idea of sparse interpolation techniques to compute reference
74 points for upscaled permeabilities. Furthermore, for an arbitrary realization, the upscaled
75 permeability is interpolated using the pre-computed values of upscaled permeabilities at ref-
76 erence points. This procedure saves the computational time by avoiding the computation
77 of upscaled quantities. Though for upscaled permeabilities such a saving may not be very
78 important, for the computations of pseudo relative permeabilities or when using global up-
79 scaling methods, these computational saving is important. We note that one can consider
80 more general ensemble level upscaling methods such as those presented in [2].

81 Two-stage MCMC methods considered here modifies the instrumental probability dis-
82 tribution by filtering the proposals via simplified models. Our numerical results show that
83 using inexpensive coarse-scale computations one can increase the acceptance rate of MCMC
84 calculations. Here the acceptance rate refers to the ratio between the number of accepted
85 permeability samples and the number of times of solving the fine-scale nonlinear PDE system.
86 In offline computational stage, we use several hundreds realizations of the permeability field
87 to construct an error model and develop ensemble level upscaling. For the error modeling,
88 we use piece-wise linear functions to fit the scattered data representing coarse-scale errors vs.
89 fine-scale errors. At the first stage, using coarse-scale runs we determine whether or not to
90 run the fine-scale simulations. To compute the approximation of the fine-scale error, offline
91 statistical models are used. If the proposal is accepted at the first-stage, then a fine-scale
92 simulation is performed at the second stage to determine the acceptance probability of the
93 proposal. The first stage of the MCMC method modifies the proposal distribution. It is easy
94 to show that the modified Markov chain satisfies the detailed balance condition for the cor-
95 rect distribution. In the paper, we also discuss the efficiency of two-stage MCMC methods.

96 We would like to note that two-stage MCMC algorithms have been used previously (e.g.,
97 [3, 19, 26, 17]) in different situations.

98 Numerical results for permeability fields generated using two-point geostatistics are pre-
99 sented in the paper. Using the Karhunen-Loève expansion, we can represent the high dimen-
100 sional permeability field by a number of parameters. Furthermore, static data (the values of
101 permeability field at some sparse locations) can be easily incorporated into the Karhunen-
102 Loève expansion to further reduce the dimension of the parameter space. Numerical results
103 are presented for black oil model (three phase flow and transport) with 8 production wells
104 and 1 injection well. In all the simulations, we observe nearly two times increase in the
105 acceptance rate. In other words, the preconditioned MCMC method can accept the same
106 number of samples with much less fine-scale runs.

107 The paper is organized in the following way. In the next section, we briefly describe
108 the model equations and their upscaling. Section 3 is devoted to the description of MCMC
109 methods. Numerical results are presented in Section 4.

110 2 Fine and coarse models

111 In this section we briefly introduce the fine- and coarse-scale models used in the simulations.
112 We consider black oil model in a subsurface formation (denoted by Ω) under the assumption
113 that the displacement is dominated by viscous effects. We neglect the effects of gravity and
114 capillary pressure, although our proposed approach is independent of the choice of physical
115 mechanisms. Porosity will be considered to be constant. The phases will be referred to as
116 water, oil and gas, designated by subscripts w , o , and g , respectively. Simultaneous flow of
117 three phases is governed by the following three equations (e.g., [5])

$$\frac{\partial}{\partial t} \left(\frac{S_j}{B_j} \right) + \nabla \cdot \left(\frac{k(x)k_{rj}}{\mu_j B_j} \nabla p_j \right) = q_j^s, \quad j = w, o, \quad (1)$$

$$\frac{\partial}{\partial t} \left(\frac{S_g}{B_g} + \frac{S_o R_{so}}{B_o} \right) + \nabla \cdot \left(\frac{k(x) k_{ro} R_{so}}{\mu_o B_o} \nabla p_o + \frac{k(x) k_{rg}}{\mu_g B_g} \nabla p_g \right) = q_g^s, \quad (2)$$

118 where B_j ($j = w, o$) is the formation volume factor of phase j , $k(x)$ is heterogeneous absolute
 119 permeability field, q^s is the source terms, p_j is the pressure of the phase j , k_{rj} is the relative
 120 permeabilities, S_j is the saturation of the phase j and R_{so} is the solubility of gas in oil.

121 Next, we will briefly describe single-phase flow upscaling procedure. These types of
 122 approaches for upscaling are discussed by many authors; see e.g., [9]. The main idea of this
 123 approach is to upscale the absolute permeability field $k(x)$ on the coarse-grid (see Figure 1),
 124 then solve the original system on the coarse-grid with upscaled permeability field. Below, we
 125 discuss briefly the upscaling of absolute permeability and ensemble level upscaling methods
 126 used in our simulations.

127 Consider the fine-scale permeability that is defined in the domain with underlying fine
 128 grid as shown in Figure 1. On the same graph we illustrate a coarse-scale partition of the
 129 domain. To calculate the upscaled permeability field at the coarse-level, we use the solutions
 130 of local pressure equations. The main idea of the calculation of a coarse-scale permeability
 131 is that it delivers the same average response as that of the underlying fine-scale problem
 132 locally. For each coarse domain K , we solve the local problems

$$\text{div}(k(x) \nabla \phi_j) = 0, \quad (3)$$

133 with some coarse-scale boundary conditions. Here $k(x)$ denotes the fine-scale permeability
 134 field. We will use the boundary conditions which are given by $\phi_j = 1$ and $\phi_j = 0$ on the
 135 opposite sides along the direction e_j and no flow boundary conditions on all other sides. For
 136 these boundary conditions, the coarse-scale permeability tensor is given by

$$(k^* e_j, e_l) = \frac{1}{|K|} \int_K (k(x) \nabla \phi_j(x), e_l) dx, \quad (4)$$

137 where ϕ_j is the solution of (3) with prescribed boundary conditions. Various boundary
 138 condition can have some influence on the accuracy of the calculations, including periodic,
 139 Dirichlet and etc. These issues have been discussed in e.g., [30]. In particular, for determining
 140 the coarse-scale permeability field one can choose the local domains that are larger than
 141 the target coarse block, K , for (3). Once the upscaled absolute permeability is computed,
 142 the original equation is solved on the coarse-grid, without changing the form of relative
 143 permeability curves.

144 Ensemble level upscaling methods compute the upscaled permeabilities based on values
 145 of upscaled permeability fields for some realizations which are computed offline. To demon-
 146 strate the concept, we assume that the permeability field is computed for realizations ω_1 ,
 147 ..., ω_N and the values are $k^*(x, \omega_1)$, ..., $k^*(x, \omega_N)$. In general, ω is infinite dimensional,
 148 though in applications considered in this paper, the permeability fields are characterized by
 149 two-point correlation functions on a fine grid and ω will be taken to be finite dimensional.
 150 Ensemble level upscaling attempts to approximate $k(x, \omega)$ for any value of ω using $k^*(x, \omega_1)$,
 151 ..., $k^*(x, \omega_N)$. First work in this direction ([2]) uses statistical approach for this approxima-
 152 tion. In this paper, we employ deterministic interpolation theory to approximate $k^*(x, \omega)$
 153 given $k^*(x, \omega_1)$, ..., $k^*(x, \omega_N)$. In our simulations, we will be using linear relations for log of
 154 permeabilities and sparse interpolation techniques in high dimensional space (e.g., [31]) to
 155 approximate $k^*(x, \omega)$. In particular,

$$\log(k^*(x, \omega)) = \sum_i L_i(\omega) \log(k^*(x, \omega_i)),$$

156 where $L_i(\omega)$ are interpolation weights which are readily available for interpolations considered
 157 here.

158 As for the quantities which will be used to condition the permeability field, we take
 159 gas/oil ratio (commonly abbreviated GOR). When oil is brought to surface conditions it is

160 usual for some gas to come out of solution. GOR is the ratio of the gas that comes out of the
 161 solution, to the volume of oil. Our goal in this paper, is to sample the fine-scale permeability
 162 field based on GOR which is a function of time in each producing well. The relation between
 163 GOR and the permeability field is highly nonlinear and can not be accurately described via
 164 linear relations.

165 3 Methodology

166 To find the permeability field given GOR information, we assume that an observed GOR,
 167 $F^{ref}(t)$, is given. Consequently, one can consider this problem as a sampling from the
 168 conditional distribution $P(k|F^{ref})$. Using Bayes theorem we can write

$$P(k|F^{ref}) \propto P(F^{ref}|k)P(k). \quad (5)$$

169 The normalizing constant in this expression is not important, because we use iterative updat-
 170 ing procedure. In (5), $P(F^{ref}|k)$ represents the likelihood function and requires the forward
 171 solution of black oil model. We will be using Metropolis-Hasting MCMC (see [27]) to sample
 172 from the posterior distribution $P(k|F)$. The main idea of MCMC is to generate a Markov
 173 chain whose stationary distribution is given by $P(k|F^{ref})$. At each iteration, a permeability
 174 field, k , is proposed using instrumental distribution $q(k|k_n)$ (where k_n is previously accepted
 175 permeability field), and then forward problem is solved to determine the acceptance proba-
 176 bility,

$$Pr(k_n, k) = \min \left(1, \frac{q(k_n|k)P(k|F^{ref})}{q(k|k_n)P(k_n|F^{ref})} \right), \quad (6)$$

177 i.e. $k_{n+1} = k$ with probability $Pr(k_n, k)$, and $k_{n+1} = k_n$ with probability $1 - Pr(k_n, k)$.

178 Since each proposal requires the fine-scale computation, direct (full) MCMC is expensive.
 179 Typically, direct MCMC requires many iterations for the convergence to a steady state, where

180 each iteration involves the computation of the fine-scale solution over a large time interval.
 181 One way to achieve efficiency is to propose an algorithm that increases the acceptance rate of
 182 MCMC. This minimizes rejection of proposals after detailed flow and transport calculations.
 183 In this paper, we use coarse-scale solutions based on single-phase upscaling to increase the
 184 acceptance rate. The main idea of this algorithm is to compare GOR that correspond to the
 185 coarse-scale models to determine whether or not to run fine-scale simulations.

186 To formulate the algorithm, we introduce several notations. For our numerical results,
 187 we will sample the likelihood

$$\pi(k) = P(F^{ref}|k) \propto \exp\left(-\frac{\|F^{ref} - F_k\|^2}{\sigma_f}\right), \quad (7)$$

188 where F^{ref} is GOR, F_k is GOR that is obtained from the simulations with permeability k ,
 189 and $\Sigma = \sigma_f I$ is covariance matrix representing the measurement errors. Here $\|F^{ref} - F_k\| =$
 190 $\left(\int_0^T |F^{ref} - F_k|^2 dt\right)^{1/2}$ is the fine-scale error between the simulated and the observed data.
 191 We note that GOR is a function of time at producing wells. In our simulations, we start
 192 with a number of realizations (usually 100-200) of the permeability field and construct the
 193 error model between $\|F_k - F^{ref}\|$ and $\|F_{k^*} - F^{ref}\|$ (see Figure 7). Using known statistical
 194 methods we derive nonlinear relation between these quantities

$$\|F_k - F^{ref}\| \approx G(\|F_{k^*} - F^{ref}\|),$$

195 where G is a nonlinear function which is estimated based on a limited number of realizations
 196 of the permeability field. G can be assumed to be random as it is done in our simulations.
 197 In our simulations, we use piece-wise Gaussian processes to fit the relation $\|F_k - F^{ref}\|$ vs.
 198 $\|F_{k^*} - F^{ref}\|$. In this case, the surrogate probability distribution used in the simulations is

$$\pi^*(k) = P(F^{ref}|k^*) \propto \exp\left(-\frac{1}{\sigma_f} \frac{G_0(\|F_{k^*} - F^{ref}\|)}{\sigma_{k^*}}\right),$$

199 where G_0 and σ_{k^*} are the mean and the variance of the piece-wise Gaussian for a given k^*
 200 (see Figure 7). The subscript k^* is used in σ to indicate that the variance of piece-wise
 201 Gaussian model is a function of k^* . We note that one can use more general probabilistic
 202 models to model error relations and also use the full covariance matrix corresponding to the
 203 measurement errors.

204 **Algorithm (two-stage MCMC with nonlinear error model)**

205 • **Offline:** Start with offline computations of GOR cross-plot by computing $\|F_k - F^{ref}\|$
 206 and $\|F_{k^*} - F^{ref}\|$. Estimate the operator G , such that $\|F_k - F^{ref}\| \approx G(\|F_{k^*} - F^{ref}\|)$.

207 • **Offline:** Compute $k^*(x, \omega_i)$ for some realizations ω_i .

208 • Step 1. At k_n generate k from $q(k|k_n)$.

209 • Step 2. Accept k for *the fine-scale run* with probability

$$g(k_n, k) = \min \left(1, \frac{q(k_n|k)\pi^*(k)}{q(k|k_n)\pi^*(k_n)} \right), \quad (8)$$

210 i.e. $k_{n+1} = k$ (conditionally) with probability $h(k_n, k)$, and $k_{n+1} = k_n$ (conditionally)
 211 with probability $1 - g(k_n, k)$. If rejected go to step 1.

212 • Step 3. Accept k with probability

$$Pr(k_n, k) = \min \left(1, \frac{Q(k_n|k)\pi(k)}{Q(k|k_n)\pi(k_n)} \right), \quad (9)$$

213 i.e. $k_{n+1} = k$ with probability $Pr(k_n, k)$, and $k_{n+1} = k_n$ with probability $1 - Pr(k_n, k)$.

214 The proposal function $Q(k|k_n)$ satisfies

$$Q(k|k_n) = g(k_n, k)q(k|k_n) + \left(1 - \int g(k_n, k)q(k|k_n)dk \right) \delta_{k_n}(k). \quad (10)$$

215 The expression for Q can be simplified (e.g., [11])

$$Pr(k_n, k) = \min \left(1, \frac{\pi(k)\pi^*(k_n)}{\pi(k_n)\pi^*(k)} \right). \quad (11)$$

216 One can also add new data to improve the estimates; however, this can introduce bias and
 217 will be avoided in our simulations.

218 In [11], it was shown that the detailed balance condition holds and MCMC converges to
 219 the correct distribution. This proof applies here. We would like to note that from (11) one
 220 obtains (see [10, 11]) that

$$Pr(k_n, k) > \exp \left(-\frac{\|F_k - F^{ref}\| - \frac{G_0(\|F_{k^*} - F^{ref}\|)}{\sigma_{k^*}}}{\sigma_f} + \frac{\|F_{k_n} - F^{ref}\| - \frac{G_0(\|F_{k_n^*} - F^{ref}\|)}{\sigma_{k^*}}}{\sigma_f} \right).$$

221 It is clear from this expression that if the error in $\|F_k - F^{ref}\| - G_0(\|F_{k^*} - F^{ref}\|)/\sigma_{k^*}$ is
 222 small for a generic k , the acceptance probability is close to 1. Thus, if the approximation
 223 with G is accurate and the spread of this approximation is low, then one can achieve high
 224 acceptance rate. Clearly, if the coarse-scale model approximates the results of the fine-scale
 225 simulations, then one can achieve high acceptance rate. However, in general, single-phase
 226 upscaling techniques do not provide accurate approximations of the fine-scale results for
 227 three-phase systems and thus modeling the relation between coarse- and fine-scale models
 228 are needed. The errors associated with this modeling will affect the efficiency of two-stage
 229 MCMC methods. In this paper, we explore this via numerical simulations.

230 We again note this approach extends previous approaches within two-stage MCMC meth-
 231 ods [11, 10] where linear relations between coarse- and fine-scale responses are used. In the
 232 numerical examples considered in this paper, we have found poor performance when linear
 233 models are used.

4 Numerical results

For our numerical tests, we use the Karhunen-Loève expansion (KLE) [21, 29] to obtain the permeability field in terms of an optimal L^2 basis. By truncating KLE, we can represent the permeability matrix by a small number of random parameters. To impose the hard constraints (the values of the permeability at prescribed locations), one can find a linear subspace of our parameter space (a hyperplane) which yields the corresponding values of the permeability field. First, we briefly recall the facts of the KLE. Denote $Y(x, \omega) = \log[k(x, \omega)]$, where the random element ω is included to remind us that k is a random field. We assume that $E[Y(x, \omega)] = 0$. Suppose $Y(x, \omega)$ is a second order stochastic process with $E \int_{\Omega} Y^2(x, \omega) dx < \infty$, where E is the expectation operator. Given an orthonormal basis $\{\phi_k\}$ in $L^2(\Omega)$, we can expand $Y(x, \omega)$ as

$$Y(x, \omega) = \sum_{k=1}^{\infty} Y_k(\omega) \phi_k(x), \quad Y_k(\omega) = \int_{\Omega} Y(x, \omega) \phi_k(x) dx.$$

We are interested in the special L^2 basis $\{\phi_k\}$ which makes the random variables Y_k uncorrelated. That is, $E(Y_i Y_j) = 0$ for all $i \neq j$. Denote the covariance function of Y as $R(x, y) = E[Y(x)Y(y)]$. Then such basis functions $\{\phi_k\}$ satisfy

$$E[Y_i Y_j] = \int_{\Omega} \phi_i(x) dx \int_{\Omega} R(x, y) \phi_j(y) dy = 0, \quad i \neq j.$$

235 Since $\{\phi_k\}$ is a complete basis in $L^2(\Omega)$, it follows that $\phi_k(x)$ are eigenfunctions of $R(x, y)$:

$$\int_{\Omega} R(x, y) \phi_k(y) dy = \lambda_k \phi_k(x), \quad k = 1, 2, \dots, \quad (12)$$

236 where $\lambda_k = E[Y_k^2] > 0$. Furthermore, we have

$$R(x, y) = \sum_{k=1}^{\infty} \lambda_k \phi_k(x) \phi_k(y). \quad (13)$$

237 Denote $\eta_k = Y_k / \sqrt{\lambda_k}$, then η_k satisfy $E(\eta_k) = 0$ and $E(\eta_i \eta_j) = \delta_{ij}$. It follows that

$$Y(x, \omega) = \sum_{k=1}^{\infty} \sqrt{\lambda_k} \eta_k(\omega) \phi_k(x), \quad (14)$$

238 where ϕ_k and λ_k satisfy (12). We assume that the eigenvalues λ_k are ordered as $\lambda_1 \geq \lambda_2 \geq \dots$

239 The expansion (14) is called the Karhunen-Loève expansion. In the KLE (14), the L^2 basis
 240 functions $\phi_k(x)$ are deterministic and resolve the spatial dependence of the permeability field.

241 The randomness is represented by the scalar random variables η_k .

After we discretize the domain Ω by a rectangular mesh, the continuous KLE (14) is reduced to a finite number of terms. In our paper, we work with finite dimensional covariance matrices defined over the square domain with 50×50 resolution. As a consequence, the covariance matrix is 2500×2500 . Note that we only need to keep the leading order terms (quantified by the magnitude of λ_k) and still capture most of the energy of the stochastic process $Y(x, \omega)$. For an N -term KLE approximation $Y_N = \sum_{k=1}^N \sqrt{\lambda_k} \eta_k \phi_k$, define the energy ratio of the approximation as

$$e(N) := \frac{E\|Y_N\|^2}{E\|Y\|^2} = \frac{\sum_{k=1}^N \lambda_k}{\sum_{k=1}^{\infty} \lambda_k}.$$

242 If $\lambda_k, k = 1, 2, \dots$, decay very fast, then the truncated KLE would be a good approximation
 243 of the stochastic process in the L^2 sense.

244 It is common to use variogram instead of covariance functions for stochastic permeability
 245 fields. The relation between them can be easily written as $\gamma(x, y) = C - R(x, y)$, where
 246 $C = E(Y(x, \omega)^2)$ is constant for stationary processes and $\gamma(x, y)$ denotes the variogram.

247 Typical variograms used in the modeling of subsurface processes are exponential, normal,
248 and spherical [6]. In this paper, we will use spherical variogram and denote correlation
249 lengths by l_1, l_2 and the variance of $\log(k)$ by $\sigma_{\log(k)}$. We first solve the eigenvalue problem
250 (12) numerically on the rectangular mesh and obtain the eigenpairs $\{\lambda_k, \phi_k\}$. Further, we
251 can sample $Y(x, \omega)$ from the truncated KLE (14) by generating Gaussian random variables
252 η_k .

253 In all numerical simulations, we will assume that the reservoir is filled with oil and water
254 is injected to displace the oil. We consider 9 spot pattern of water flooding. There is one
255 water injector, and 8 producers (see Figure 2). The domain is taken to be square with
256 50×50 fine grid resolution with dimensions 50×50 . The coarse grid is taken to be uniform
257 5×5 in all cases. As for ensemble level upscaling, we use lowest order Smolyak interpolation
258 which provides $2d + 1$ nodes in d -dimensional space. We refer to [31] for the description of
259 Smolyak interpolation and to [7] the results on upscaling using Smolyak interpolation. As
260 we mentioned earlier that ensemble level upscaling techniques used here simply interpolate
261 the log of upscaled permeability based on the values of log of upscaled permeability at some
262 regularly spaced points in high dimensions. Because the points are regularly spaced the
263 interpolation formula and interpolation weights can be easily derived (see [31]).

264 Solution gas/oil ratio and gas formation volume factor are shown in Figure 3 (left) and
265 Figure 3 (right). From right top figure, the bubble point pressure of the reservoir is 3000
266 psia. Relative permeability of water, oil, and gas are shown in Figure 4 (left) and in Figure
267 4 (right). Modified Stones II second three-phase relative permeability model was used to
268 compute oil relative permeability [1]. We note that if reservoir pressure is above the bubble
269 point pressure, the flow is two-phase (water and oil); if the pressure drops below the bubble
270 point pressure, then the gas evolves into a liquid phase and a gas phase. The flow is three-
271 phases: water, oil and gas. There are maximum of three components, water, oil and gas.
272 In the black oil model, it is assumed that no mass transfer occurs between the water phase

273 and other two phases. Moreover, the mass fractions of the oil and gas components in the oil
274 phase can be determined by gas solubility.

275 In the first example, we use a reference permeability field with the correlation lengths
276 $l_x = 30$, $l_y = 2$, and the variance of $\log(k)$ is 2. A realization of this permeability field is
277 used to generate a reference permeability field. In the sampling procedure, we choose the
278 permeability fields with different correlation lengths (still on 50×50 fine grid resolution).
279 In particular, we choose $l_x = 18$, $l_y = 3$ and keep only 50 eigenvalues/eigenvectors in KLE.
280 In Figure 5, GOR misfit vs. iterations are plotted. We compare the use of two coarse scale
281 models. In the first case, we use the coarse-scale model where no approximation is made in
282 computing upscaled permeabilities. In the second example, we use ensemble level upscaled
283 permeability which is computed based on offline computations. We note that the second
284 case is less accurate though less expensive to compute. We observe from this figure that
285 all three curves show that the two-stage MCMC methods have similar convergence as the
286 full MCMC. These results are based on 1000 proposals. The acceptance rates are given
287 as following, full MCMC - 8.8 %, two-stage MCMC with upscaling is 22 %, and two-stage
288 MCMC with ensemble level upscaling is 17 %. We observe that two-stage MCMC methods
289 provide nearly two times higher acceptance rates. Since the computational cost associated
290 with coarse-scale simulations is negligible, two-stage MCMC methods improve CPU required
291 for sampling the posterior distribution by nearly two-fold. In all simulations, random walk
292 instrumental distribution for $q(k|k_n)$ is used, where $k = k_n + \epsilon_n$ with ϵ_n being a random
293 perturbation with prescribed distribution. The formal convergence diagnosis can be performed
294 using multiple chains method based convergence diagnosis ([14]). In this paper, our goal is
295 to compare modified chain with the chain obtained via direct MCMC, and thus we restrict
296 ourselves to only showing RMS vs. the number of iterations. We note that the convergence
297 diagnostics has nothing to do with the rate of convergence, which depends on the second
298 largest eigenvalue of the transition matrix of the Markov chain. Figure 6 shows the GOR

299 matches of the producers. As we see that the sampled realizations match GOR very well.
300 In these figures, reference GOR is designated by green color, and initial GOR is designated
301 by blue, and the sampled GORs are designated by red color. The data which is used to
302 model the error is presented in Figure 7. We plot both mean as well as mean plus/minus
303 standard deviation. In our simulations, we have used piecewise linear relation to model the
304 mean behavior and constant variation within each bin. Furthermore, Gaussian distribution
305 is used for each bin. In Figure 8, the samples of the permeability field are depicted. One
306 can observe that the initial model does not have the high flow channel located at upper side
307 of the domain, while the corrected permeability models have this high conductivity region.

308 In the second example, the reference permeability field is chosen with the correlation
309 lengths $l_x = 25$ and $l_y = 6$ and 100 eigenvalues/eigenvectors are kept in KLE. For sampling
310 purposes, the permeability fields are generated with 50×50 fine resolution and with $l_x = 20$
311 and $l_z = 4$ and 40 eigenvalues/eigenvectors are kept in KLE. In Figure 9, GOR misfit vs.
312 iterations are plotted. As before, we compare the use of two coarse scale models. In the
313 first case, we use coarse-scale model where no approximation is made in computing upscaled
314 permeabilities. In the second example, we use ensemble level upscaled permeability which is
315 computed based on offline computations. We observe from this figure that all three curves
316 show that the two-stage MCMC has similar convergence as the full MCMC. These results
317 are based on 1000 proposals. The acceptance rates are given as following, full MCMC -
318 12 %, two-stage MCMC with upscaling is 24 %, and two-stage MCMC with ensemble level
319 upscaling is 20 %. Figure 10 shows the GOR matches of the producers. As we see that
320 the sampled realizations match GOR very well. As before, reference GOR is designated by
321 green color, and initial GOR is designated by blue, and the sampled GORs are designated
322 by red color. The data which is used to model the error is presented in Figure 11. In our
323 simulations, we have used piecewise linear relation to model the mean behavior and constant
324 variation within each segment.

5 Conclusions

In this paper, we study two-stage MCMC methods which use simplified and inexpensive upscaled models to speed-up the sampling of the posterior distribution. Our underlying equations describe the flow and transport of three-phase flow system in heterogeneous porous media. We employ single-phase upscaling methods for coarsening the flow and transport equations. Because single-phase upscaling techniques are not very accurate for complex black oil models (see e.g., Figure 11), the proposed algorithm requires offline computations where a nonlinear relation between the coarse- and fine-scale models are constructed. This relation is used within the context of two-stage MCMC to perform rigorous sampling. The proposed method generalizes the existing multi-stage MCMC methods where a linear relation between coarse- and fine-scale models are used. In our coarse-scale models, we also use ensemble level upscaling techniques for inexpensive coarse-scale computations. The efficiency of two-stage MCMC methods depends on the accuracy of the nonlinear approximation and the spread in this approximation. The modeling errors in this approximation affect the efficiency of two-stage MCMC methods. In this paper, we explore this via numerical simulations. Numerical results study the advantages, efficiency and utility of the method for uncertainty assessment in the history matching. We show that two-stage MCMC methods provide two times higher acceptance rates, and thus improve CPU required for sampling the posterior distribution by nearly two-fold.

References

- [1] K. AZIZ, *Petroleum reservoir simulation*, Chapman & Hall, 1979.
- [2] Y. CHEN AND L. DURLOFSKY, *An ensemble level upscaling approach for efficient estimation of fine-scale production statistics using coarse-scale simulations*, SPE paper

- 106086, presented at the SPE Reservoir Simulation Symposium, Houston, Feb. 26-28 (2007).
- [3] A. CHRISTEN AND C. FOX, *MCMC using an approximation*, Technical report, Department of Mathematics, The University of Auckland, New Zealand.
- [4] M. CHRISTIE, J. GLIMM, J. GRIVE, D. HIGHON, D. SHARP, AND M. WOOD-SCHULTZ, *Error analysis and simulations of complex phenomena*, Los Alamos Sci. 29 (2005) 625.
- [5] A. DATTA-GUPTA AND M.J. KING, *Streamline Simulation: Theory and Practice*, Publisher: Society of Petroleum Engineers, 2007
- [6] C. V. DEUTSCH AND A. G. JOURNEL, *GSLIB: Geostatistical software library and user's guide, 2nd edition*, Oxford University Press, New York, 1998.
- [7] P. DOSTERT, *Uncertainty Quantification Using Multiscale Methods for Porous Media Flows*, Ph.D. thesis, Texas A&M University, 2007
- [8] P. DOSTERT, Y. EFENDIEV, T. HOU, AND W. LUO, *Coarse-gradient Langevin algorithms for dynamic data integration and uncertainty quantification*, 217 (1), pp.123-142, 2006
- [9] L. J. DURLOFSKY, *Numerical calculation of equivalent grid block permeability tensors for heterogeneous porous media*, Water Resour. Res., 27 (1991), pp. 699–708.
- [10] Y. EFENDIEV, A. DATTA-GUPTA, V. GINTING, X. MA, AND B. MALLICK, *An efficient two-stage Markov chain Monte Carlo method for dynamic data integration*, 41, W12423, doi:=10.1029/2004WR003764
- [11] Y. EFENDIEV, T. HOU AND W. LUO, *Preconditioning Markov chain Monte Carlo simulations using coarse-scale models*, SIAM. Sci. Comp. 28(2), pp. 776-803.

- [12] V. GINTING, *Analysis of two-scale finite volume element method for elliptic problem*, Journal of Numerical Mathematics, 12(2) (2004), pp. 119–142.
- [13] P. FRAUENFELDER, C. SCHWAB AND R. A. TODOR, *Finite elements for elliptic problems with stochastic coefficients*, Comput. Methods Appl. Mech. Engrg., 194 (2005), pp. 205–228.
- [14] Gelman, A., and Rubin, B., 1992, *Inference from iterative simulation using multiple sequences*, Statistical Science, 7, pp. 457-511.
- [15] J. GLIMM AND D. H. SHARP, *Prediction and the quantification of uncertainty*, Phys. D, 133 (1999), pp. 152–170. Predictability: quantifying uncertainty in models of complex phenomena (Los Alamos, NM, 1998).
- [16] J. GLIMM, S. HOU, Y.H. YEE, D.H. SHARP AND K. YE, *Solution error models for uncertainty quantification*, AMS Advances in Differential Equations and Mathematical Physics (2002).
- [17] D. HIGDON, H. LEE AND Z. BI, *A Bayesian approach to characterizing uncertainty in inverse problems using coarse and fine-scale information*, IEEE Transactions on Signal Processing, 50(2) (2002), pp. 388-399.
- [18] P. KITANIDIS, *Quasi-linear geostatistical theory for inversing*, Water Resour. Res., 31 (1995), pp. 2411–2419.
- [19] J. S. Liu, *Monte Carlo Strategies in Scientific Computing*, Springer, New-York, 2001.
- [20] O. LODOEN, H. OMRE, L. DURLOFSKY, AND Y. CHEN, *Assessment of uncertainty in reservoir production forecasts using upscaled models*, in proceedings of the 7th International Geostatistics Congress, Banff, Canada, September 26 - October 1, 2004.
- [21] M. LOEVE, *Probability Theory*, 4th ed., Springer, Berlin, 1977.

- [22] X. MA, M. AL-HARBI, A. DATTA-GUPTA, AND Y. EFENDIEV *An Efficient Two-Stage Sampling Method for Uncertainty Quantification in History Matching Geological Models*, to appear in SPEJ.
- [23] S. P. MEYN, R. L. TWEEDIE, *Markov Chains and Stochastic Stability*, Springer-Verlag, London, 1996.
- [24] D. OLIVER, L. CUNHA, AND A. REYNOLDS, *Markov chain Monte Carlo methods for conditioning a permeability field to pressure data*, *Mathematical Geology*, 29 (1997).
- [25] D. OLIVER, N. HE, AND A. REYNOLDS, *Conditioning permeability fields to pressure data*, 5th European conference on the mathematics of oil recovery, Leoben, Austria, 3-6 September, 1996.
- [26] H. OMRE AND O. P. LODOEN, *Improved production forecasts and history matching using approximate fluid flow simulators*, *SPE Journal*, September 2004, pp. 339-351.
- [27] C. ROBERT AND G. CASELLA, *Monte Carlo Statistical Methods*, Springer-Verlag, New-York, 1999.
- [28] A. O'SULLIVAN, *Modelling Simulation Error For Improved Reservoir Prediction*, PhD Thesis, (Heriot Watt University, 2004).
- [29] E. WONG, *Stochastic Processes in Information and Dynamical Systems*, McGraw-Hill, 1971.
- [30] X. H. WU, Y. EFENDIEV, AND T. Y. HOU, *Analysis of upscaling absolute permeability*, *Discrete and Continuous Dynamical Systems, Series B*, 2 (2002), pp. 185–204.
- [31] D. XIU AND J. HESTHAVEN, *High-Order Collocation Methods for Differential Equations with Random Inputs*, *SIAM J. Sci. Comput.* Wol. 27, No. 3, 2007, pp. 1118-1139.

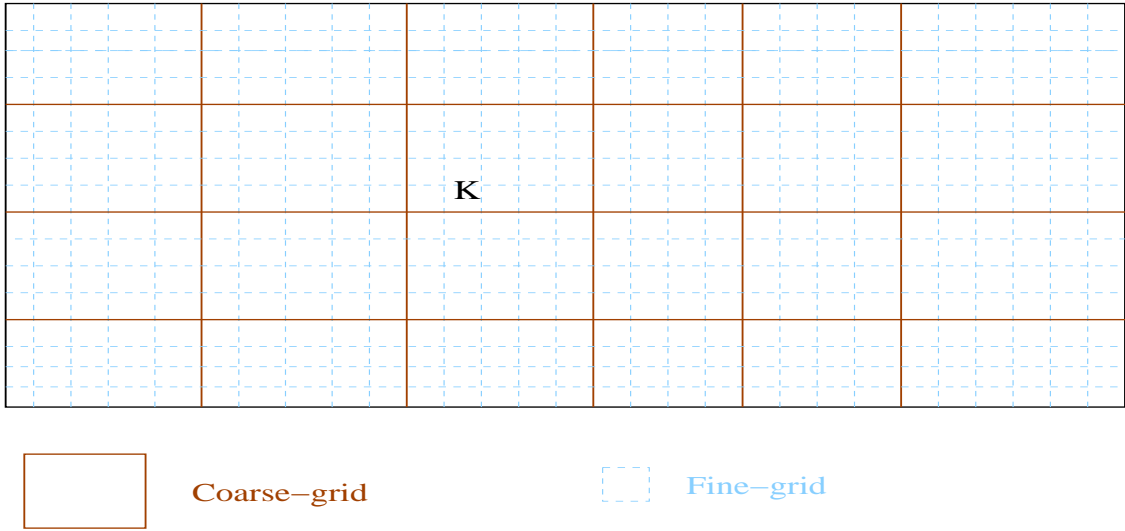


Figure 1: Schematic description of fine- and coarse-grids. Bold lines illustrate a coarse-scale partitioning, while thin lines show a fine-scale partitioning within coarse-grid cells.

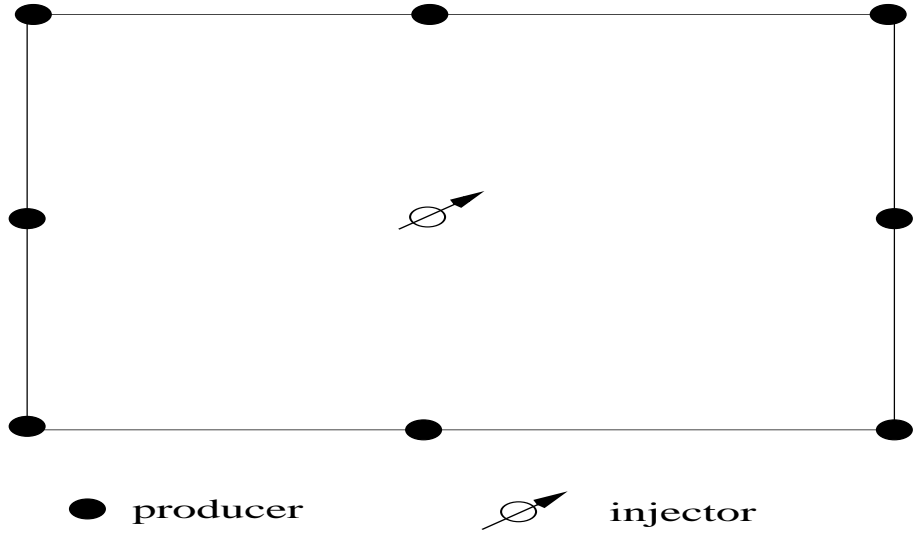


Figure 2: Well configuration used in the simulations.

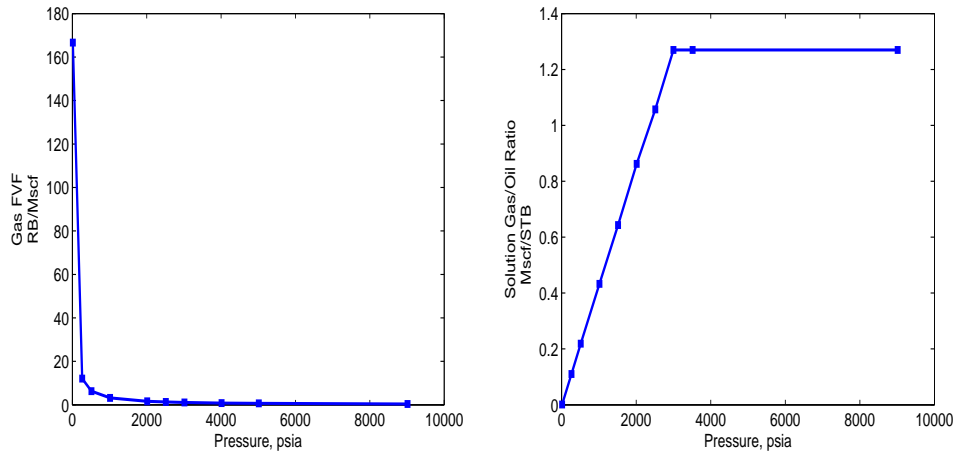


Figure 3: Solution gas/oil ratio and gas formation volume factor.

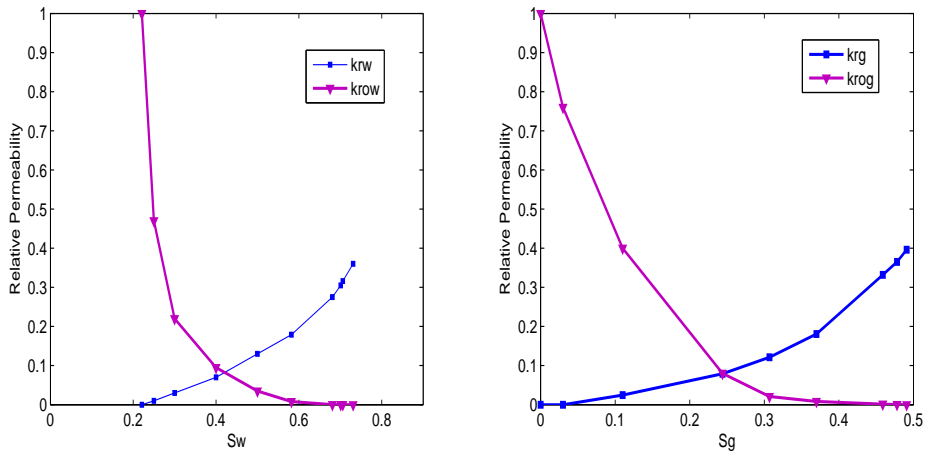


Figure 4: Relative permeabilities used in the simulations.

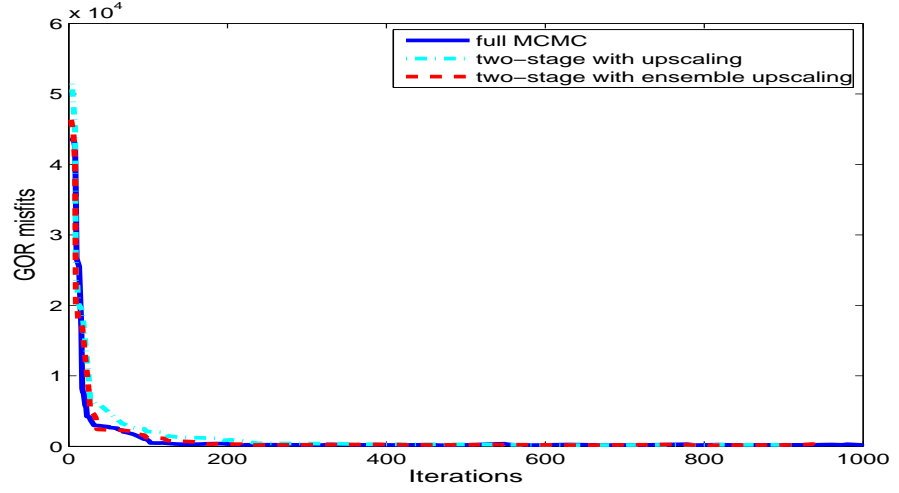


Figure 5: Error vs. proposal iterations.

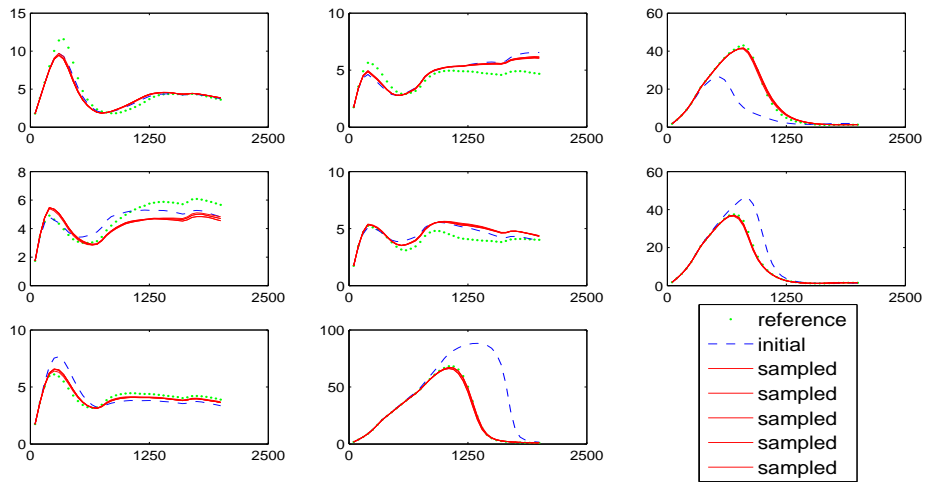


Figure 6: GOR data at producers.

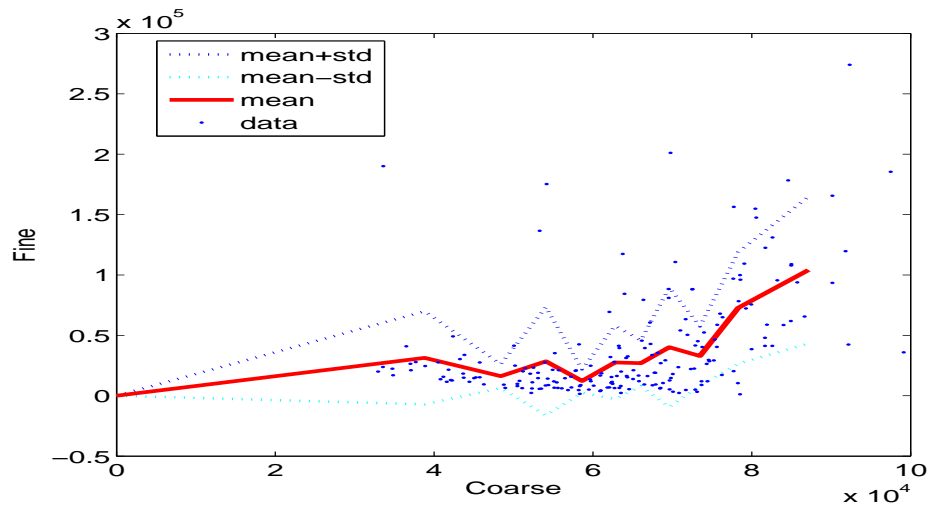


Figure 7: Error model.

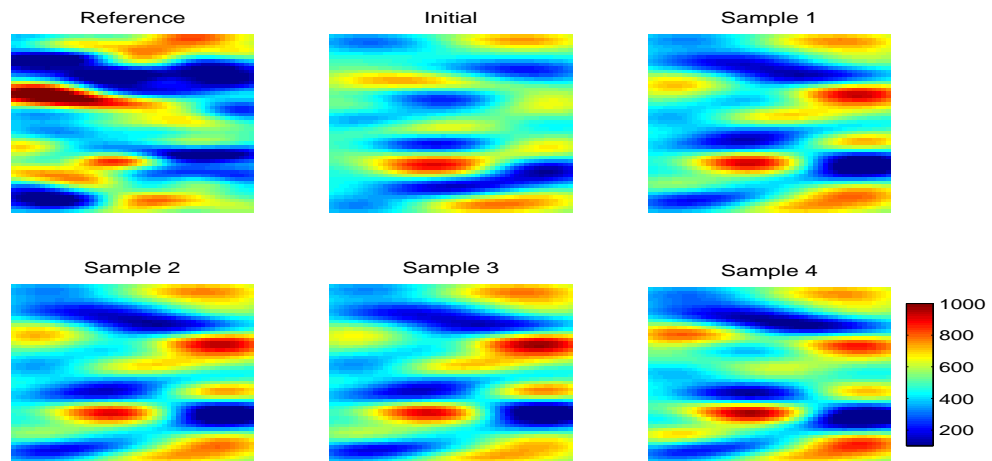


Figure 8: Permeability realizations.

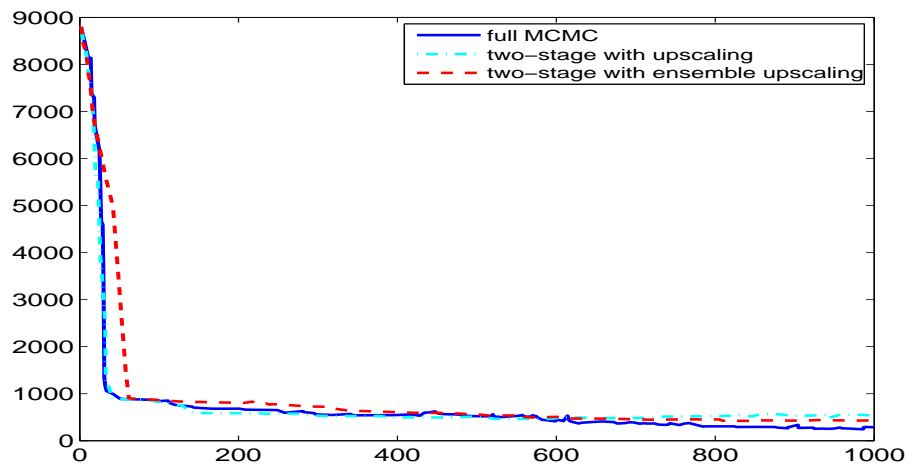


Figure 9: Error vs. proposal iterations.

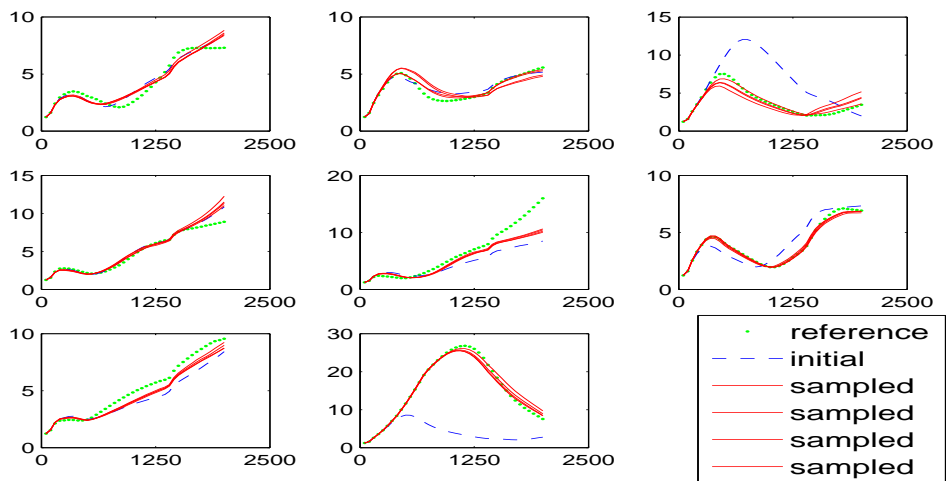


Figure 10: GOR data at producers.

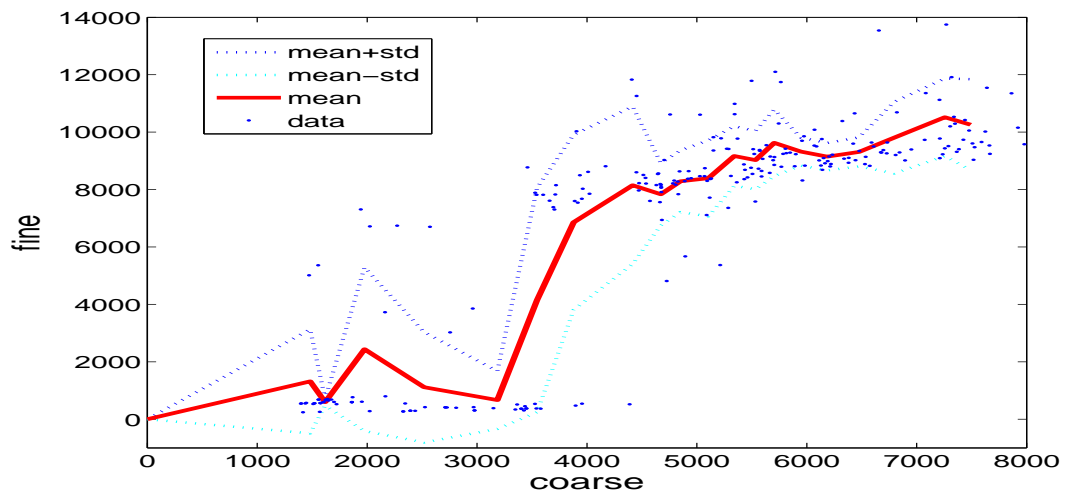


Figure 11: Error model.